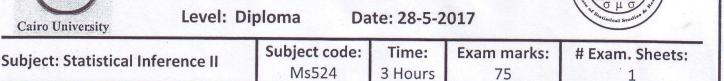


Department: Mathematical Statistics

Academic Year: 2016-2017

Academic Semester: Second



Examiner : Dr. Hiba Zeyada

Question One:

a) Neyman –Pearson lemma states that:

" If C is a critical region of size α and K is a constant such that $\frac{L_0}{L_1} \leq K$ inside C and $\frac{L_0}{L_1} \geq K$ outside C, then C is the most powerful critical region for testing $\theta = \theta_0$ against $\theta = \theta_1$ ". **Prove the above lemma**

b) Given a random sample of size n=20 from N(μ , 5) use Neyman –Pearson lemma to find the MPT of size $\alpha = 0.05$ to test H₀: $\mu = 7$ versus H₁: $\mu > 7$.

Question Two:

a) Suppose we wish to test $H_0: \mu = 15$ versus $H_1: \mu = 16$

with $\alpha = \beta = 0.05$. Find the sample size that will ensure this accuracy assume $\sigma^2 = 9$.

b) Given a random sample of size n from a normal distribution with mean μ (unknown) and variance σ^2 . For testing $H_0: \sigma^2 = \sigma_0^2$ versus $H_1: \sigma^2 > \sigma_0^2$, show that the likelihood ratio test is equivalent to the chi-square test.

<u>Question Three:</u> True or false? Correct the false

- 1- The random variable $\frac{(n-1)S^2}{\sigma^2}$ is distributed as chi-square with n-1 degrees of freedom
- 2- The p-value of a test is the largest value of α that would lead to accept the alternative hypothesis.
- 3- The random variable $\frac{\bar{x}-\mu}{s/\sqrt{n}}$ is standard normal distributed.
- 4- The power function is the probability to accept the null hypothesis
- 5- If X is a continuous random variable then the distribution of $F_X(x)$ is uniform with parameters a and b
- 6- A type I error is made if H_0 is rejected when H_1 is true
- 7- A Test is MPT of size α if it has the size of its type I error equal to α and has smallest type II error among all other tests with size of type I error α or less.

Hint: use $Z_{0.05} = 1.64$

Year:2011-2012 (Spring semester) Final Exam Time allowed: 3 hrs Marks 70

Solve the Following Ouestions

Ouestion one: (15 Marks)

- a) State and prove Nayman -Pearson theorem for the best critical region of size α for testing the simple hypothesis $H_0: \theta = \theta'$ against the alternative simple hypothesis $H_a: \theta = \theta'$
- b) Suppose that x_1, x_2, \dots, x_n denote a random sample from a population having an exponential distribution with density $f(x) = \frac{1}{\theta}e^{-\frac{x}{\theta}}$, x > 0 Derive a most powerful test

for testing $H_0: \theta = \theta_0$ vs $H_a: \theta = \theta_a$ when $\theta_a < \theta_0$.

Ouestion two: (12 Marks)

Suppose that y_1, y_2, \dots, y_n denote a random sample from a population having a Poisson distribution with probability distribution $f(x;\lambda) = \frac{e^{-\lambda}\lambda^x}{x!}$, $x = 0, 1, 2, \cdots$ find the form of the rejection region (R.R) for a most powerful test for testing $H_0: \lambda = \lambda_0$ vs $H_a: \lambda = \lambda_a$ when $\lambda_a > \lambda_0$.

Question three: (13 marks)

The number of computer malfunctions per day is recorded for 260 days. Test number of malfunctions follows Poisson distribution with parameter $\lambda = 1.15$

Number of malfunctions(x_i)	0	1	2	3	4	5
Number of days f_i	77	90	55	30	5	3

Question four: (15 marks)

distribution with mean μ and variance $\sigma^2 = 100, n = 64$ and X has a normal Let $RR = \{(x_1, x_2, \dots, x_n) | x < 98 \text{ or } x > 102\}$ and the test is $H_0: \mu = 100 \quad vs \quad H_a: \mu \neq 100$ Find: iii) power of the test ii) Type II error when $\mu = 101$ i) Type I error.

Question five: (15 marks)

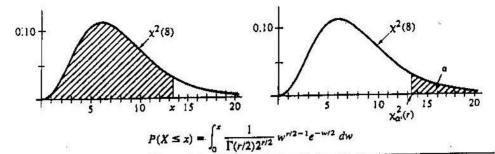
 x_1, x_2, \dots, x_n be iid random sample distributed gamma with density defined by Let

$$f(x) = \frac{x}{\theta^2} e^{-\theta}, \quad x > 0$$

- a) Verify that $\hat{\theta}$ is an unbiased estimator of θ .
- b) Compute the Cramer Rao lower bound for $\hat{\theta}$.
- c) Show that $\hat{\theta}$ is a minimum variance unbiased estimator.
- d) Is $\hat{\theta}$ complete sufficient statistics.





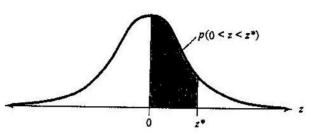


	0.010	0.025	0.050	P(X	≤ x)		0.975	0.990
				0.100	0.900	0.950		
r	x8.99(r)	X ² .975(r)	X0.95(r)	χ ² 0.90(r)	$\chi^2_{0.10}(r)$	X2.05(r)	X0.025(r)	x2.01(r)
1	0.000	0.001	0.004	0.016	2.706	3.841	5.024	6.635
2	0.000	0.051	0.103	0.211	4.605	5.991	. 7.378	9.210
2	0.115	0.216	0.352	0.584	6.251	7.815	9,348	11.34
3 4	0.297	0.484	0.711	1.064	7.779	9.488	11.14	13.28
5	0.554	0.831	1.145	1.610	9.236	11.07	12.83	15.09
6	0.872	1.237	1.635	2.204	10.64	12.59	14.45	16.81
7	1.239	1,690	2.167	2.833	12.02	14.07	16.01	18.48
8	1.646	2.180	2.733	3.490	13.36	15.51	17.54	20.09
9	2.088	2.700	3.325	4.168	14.68	16.92	19.02	21.67
10	2.558	3.247	3.940	4.865	15.99	18.31	20.48	23.21
11	3.053	3.816	4.575	5.578	17.28	19.68	21.92	24.72
	3.571	4.404	5.226	6.304	18.55	21.03	23.34	26.22
12 13	4.107	5.009	5.892	7.042	19.81	22.36	24.74	27.69
14	4.660	5.629	6.571	7.790	21.06	23.68	26.12	29.14
15	5.229	6.262	7.261	8.547	22.31	25.00	27.49	30.58
		6.908	7.962	9.312	23.54	26.30	28,84	32.00
16	5.812	7.564	8.672	10.08	24.77	27.59	30.19	33.41
17	6.408 7.015	8.231	9.390	10.86	25.99	28.87	31.53	34.80
18	7.633	8.907	10.12	11.65	27.20	30.14	32.85	36.19
19 20	8.260	9.591	10.85	12.44	28.41	31.41	34.17	37.57
		10.28	11.59	13.24	29.62	32.67	35.48	38.93
21	8.897	10.28	12.34	14.04	30.81	33.92	36.78	40.29
22	9.542	11.69	13.09	14.85	32.01	35.17	38.08	41.64
23 24	10.20	12.40	13.85	15.66	33.20	36.42	39.36	42.98
25	11.52	13.12	14.61	16.47	34.38	37.65	40.65	44.31
	and the second second		15.38	17.29	35.56	38.88	41.92	45.64
26	12.20	13.84	16.15	18,11	36.74	40.11	43.19	46.96
27	12.88	14.57	16.93	18.94	37.92	41.34	44.46	48.28
28	13.56	15.31 16.05	17.71	19.77	39.09	42.56	45.72	49.59
29	14.26	16.05	18.49	20.60	40.26	43.77	46.98.	50.89
30	14.95				51.80	55.76	59.34	63.69
40	22.16	24.43	26.51	29.05	63.17	67.50	71.42	76.15
50	29.71	32.36	34,76	37.69	74.40	79.08	83.30	88.38
60	37.48	40,48	43.19	46.46	85.53	90.53	95.02	100.4
70	45.44	48.76	51.74	55.33	96.58	101.9	106.6	112.3
80	53.34	57.15	60.39	64.28	90.00	101.7	100.0	

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F Body Table for the Standard Normal Distribution



. z*	0.00	0.01	0.02	0.03	0.04	0.05	0.06	0.07	0.08	0.09
0.0	0.0000	0.0040	0.0080	0.0120	0.0160	0.0199	0.0239	0.0279	0.0319	0.0359
0.1	0.0398	0.0438	0.0478	0.0517	0.0557	0.0596	0.0636	0.0675	0.0319	0.0339
0.2	0.0793	0.0832	0.0871	0.0910	0.0948	0.0987	0.1026	0.1064	0.1103	
0.3	0.1179	0.1217	0.1255	0.1293	0.1331	0.1368	0.1406	0.1443	0.1480	0.1141 0.1517
0.4	0.1554	0.1591	0.1628	0.1664	0.1700	0.1736	0.1772	0.1808	0.1430	
0.5	0.1915	0.1950	0.1985	0.2019	0.2054	0.2088	0.2123	0.2157	0.1844	0.1879 0.2224
0.6	0.2257	0.2291	0.2324	0.2357	0.2389	0,2422	0.2454	0.2486	0.2517	0.2549
0.7	0.2580	0.2611	0.2642	0.2673	0.2704	0.2734	0.2764	0.2794	0.2823	0.2349
0.8	0.2881	0.2910	0.2939	0.2967	0.2995	0.3023	0,3051	.0.3078	0.2823	0.2832
0.9	0.3159	0.3186	0.3212	0.3238	0.3264	0.3289	0.3315	0.3340	0.3365	0.3389
1.0	0,3413	0.3438	0.3461	0.3485	0.3508	0.3531	0.3554	0.3577	0.3599	0.3621
1.1	0.3643	0.3665	0.3686	0.3708	0.3729	0.3749	0.3770	0.3790	0.3810	0.3830
1.2	0.3849	0.3869	0.3888	0.3907	0.3925	0.3944	0.3962	0.3980	0.3810	0.3830
1.3	0.4032	0.4049	0.4066	0.4082	0.4099	0.4115	0.4131	0.4147	0.3997	0.4013
1.4	0.4192	0.4207	0.4222	0.4236	0.4251	0.4265	0.4279	0.4292	0.4102	0.4177
1.5	0.4332	0.4345	0.4357	0.4370	0.4382	0.4394	0.4406	0.4418	0.4300	0.4319
1.6	0.4452	0.4463	0.4474	0.4484	0.4495	0.4505	0.4515	0.4525	0.4535	0.4545
1.7	0.4554	0.4564	0.4573	0.4582	0.4591	0.4599	0.4608	0.4616	0.4625	0.4543
1.8	0.4641	0.4649	0.4656	0.4664	0.4671	0.4678	0,4686	0.4692	0.4699	0.4633
1.9	0.4713	0.4719	0.4726	0.4732	0.4738	0,4744	0.4750	0.4756	0.4099	0.4767
2.0	0.4772	0.4778	0.4783	0.4788	0.4793	0.4798	0.4803	0.4808	0.4812	0.4787
2.1	0.4821	0.4826	0.4830	0.4834	0.4838	0.4842	0.4846	0.4850	0.4854	0.4857
2.2	0.4861	0.4864	0.4868	0.4871	0.4875	0.4878	0.4881	0.4884	0.4887	0.4890
2.3	0.4893	0.4896	0.4898	0.4901	0.4904	0.4906	0.4909	0.4911	0.4913	0.4916
2.4	0.4918	0.4920	0.4922	0.4925	0.4927	0.4929	0.4931	0.4932	0.4913	0.4916
2.5	0.4938	0.4940	0.4941	0.4943	0.4945	0.4946	0.4948	0.4949	0.4951	0.4952
2.6	0.4953	0.4955	0,4956	0.4957	0.4959	0.4960	0.4961	0.4962	0.4963	0.4964
2.7	0.4965	0.4966	0.4967	0.4968	0.4969	0.4970	0.4971	0.4902		1
2.8	0.4974	0.4975	0.4976	0.4977	0.4977	0.4978	0.4979	0.4972	0.4973	0.4974
2.9	0.4981	0.4982	0.4982	0.4983	0.4984	0.4984	0.4985	0.4979	0.4980	0.4981
3.0	0.4987	0.4987	0.4987	0.4988	0.4988	0.4989	0.4989	0.4985	0.4986	0.4986 0.4990

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